

Bayesian Approaches for Benefit-Risk Assessment with Examples

Ram Tiwari and Chul Ahn
Division of Biostatistics
Center for Device and Radiological Health
Food and Drug Administration

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Acknowledgements

This presentation is based on collaboration with:

- Shiqi Cui, Graduate Student, University of Missouri
- Yueqin Zhao, PhD: Mathematical Statistician, DB-7, OB, CDER
- Jyoti Zalkikar, PhD: Team Leader, DB-5, OB, CDER
- Lisa LaVange, PhD: Director, OB, CDER

Papers:

- Zhao, Y., Zalkikar, J., Tiwari, R. C., and LaVange, L. M. (2014), A Bayesian approach for benefit-risk assessment, Statistics in Biopharmaceutical Research, online.
- Zhao, Y., Cui, s., and Tiwari, R.C. (2016), Bayesian approach to personalized benefit-risk assessment, Statistics in Biopharmaceutical Research, online



Outline of the talk

- Introduction
 - Some commonly-used Benefit-risk (BR) measures
- Data Format: Benefit-Risk Categories
 - Hydromorphone clinical trial data
- BR measures based on Global benefit-risk scores
- Bayesian approach based on multinomial-Dirichlet conjugate prior
 - Sequential updating
 - Power Prior
 - LPML model selection
- Bayesian approach based on Dirichlet process
- Conclusion



INTRODUCTION

- The benefit-risk assessment is the basis of regulatory decisions in the pre-market and post-market review processes.
- The evaluation of benefit and risk faces several challenges.
 - Benefit and risk are usually not measured on the same scales;
 - Patients withdrawal are not rare and usually due to many reasons;
 - Benefit and risk are not independent;
 - The benefit/risk trade-off may change over the course of a clinical trial.





- Benefit-risk ratio
- Time without symptoms of disease and toxic effects (TWiST)
- Ratio of number needed to treat for benefits (NNT(B)) to number needed to treat for risks (NNT(R))
- Multiple criteria decision analysis (MCDA)
- Integrated benefit-risk measure (MDIC)
- Global benefit-risk scores -- Chuang-Stein et al. (1991)



DATA FORMAT: BR CATEGORIES

• Chuang-Stein et al. (1991): the outcomes of a clinical trial can be classified into five mutually exclusive categories.

Table: Outcomes of a clinical trial with binary response data.

	Benefit	No benefit				
No adverse event	Category 1	Category 3				
Adverse event	Category 2	Category 4				
Withdrawal	Category 5					



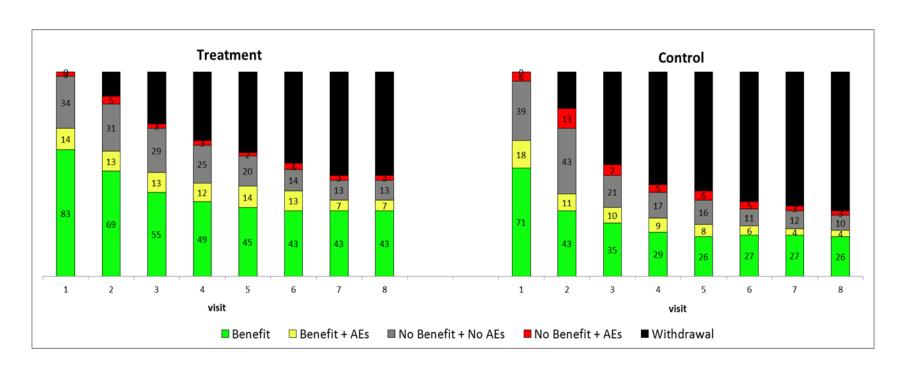
SUBJECT-LEVEL OUTCOMES OF HYDROMORPHONE TRIAL DATA

• 2 arms (treatment and control), 5 categories

134 individuals at 8 visits on each arm



EXAMPLE 1: HYDROMORPHONE DATA*



^{*}Data was provided by Jonathan Norton.



SUBJECT-LEVEL RESPONSE ON TREATMENT ARM

Subject	Visit								
	1	2	3	4	5	6	7	8	
1	1	1	1	1	1	1	1	1	
2	1	1	1	1	1	2	5	5	
134	4	5	5	5	5	5	5	5	



AGGREGATED DATA BY VISITS ON TREATMENT ARM

Category	Visit							
	1	2	3	4	5	6	7	8
1	82	65	58	50	48	43	43	44
2	15	12	10	11	15	15	8	7
3	25	30	23	20	20	15	15	17
4	12	5	3	4	2	4	3	3
5	0	22	39	49	49	57	65	63



CHUANG-STEIN ET AL. (1991): GLOBAL BR SCORES

$$BRScore_Linear = \sum_{i=1}^{2} w_{i} p_{i} - \sum_{i=3}^{5} w_{i} p_{i}$$

$$BRScore_Ratio = \frac{(\sum_{i=1}^{2} w_{i} p_{i})^{e}}{\sum_{i=3}^{5} w_{i} p_{i}}$$

$$BRScore_Cmp_Ratio = \frac{w_{1} p_{1}}{w_{5} p_{5}} (\frac{w_{2} p_{2}}{w_{3} p_{3} + w_{4} p_{4}})^{f}$$

$$where w_{1} = 2, w_{2} = 1, w_{3} = 0, w_{4} = 1, w_{5} = 2$$





 BR measures based on the global scores proposed by Chuang-Stein et al. (1991)

$$BR_Linear = (\sum_{i=1}^{2} w_{i} p_{i,T} - \sum_{i=3}^{4} w_{i} p_{i,T}) - (\sum_{i=1}^{2} w_{i} p_{i,C} - \sum_{i=3}^{4} w_{i} p_{i,C})$$

$$BR_Ratio1 = \log(\frac{(\sum_{i=1}^{2} w_{i} p_{i,T})^{e}}{\sum_{i=3}^{4} w_{i} p_{i,T}}) - \log(\frac{(\sum_{i=1}^{2} w_{i} p_{i,C})^{e}}{\sum_{i=3}^{4} w_{i} p_{i,C}})$$

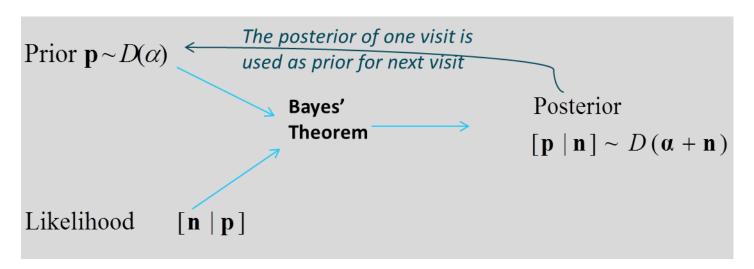
$$BR_Ratio2 = \log(\frac{(\sum_{i=1}^{2} w_{i} p_{i,T})^{e}}{w_{2} p_{2,T} + w_{4} p_{4,T}}) - \log(\frac{(\sum_{i=1}^{2} w_{i} p_{i,C})^{e}}{w_{2} p_{2,C} + w_{4} p_{4,C}})$$

 BR measures based on the global scores are for each arm (treatment and comparator) separately.



BAYESIAN APPROACH: USE OF DIRICHLET DISTRIBUTION AS A CONJUGATE PRIOR

 Dirichlet distribution is used as the conjugate prior for multinomial distribution, and the posterior distribution of the five-category random variable is derived at each visit using sequentially updated posterior as a prior.





SEQUENTIAL UPDATING

Sequential updating of the posteriors are given by:

```
Visit 1: [\mathbf{p} \mid \mathbf{n}_1] \propto [\mathbf{n}_1 \mid \mathbf{p}]^*[\mathbf{p}]

Visit 2: [\mathbf{p} \mid \mathbf{n}_1, \mathbf{n}_2] \propto [\mathbf{n}_2 \mid \mathbf{p}]^*[\mathbf{p} \mid \mathbf{n}_1]

.....

Visit M: [\mathbf{p} \mid \mathbf{n}_1, \mathbf{n}_2, ..., \mathbf{n}_M] \propto [\mathbf{n}_M \mid \mathbf{p}]^*[\mathbf{p} \mid \mathbf{n}_1, \mathbf{n}_2, ..., \mathbf{n}_{M-1}]
```

 The posterior mean (i.e., Bayes estimate) and 95% credible interval for each of the four measures are obtained using a Markov chain Monte Carlo (MCMC) technique.



DECISION RULES

- For a BR measure,
 - ➤ If the credible interval include the value zero, the benefit does not outweigh the risk;
 - ➤ If the lower bound of the credible interval is greater than zero, the benefit outweighs the risk;
 - ➤ If the upper bound of the credible interval is less than zero, the risk outweighs the benefit.



USE OF POWER PRIOR

• Power prior (Ibrahim and Chen, 2000) is used through the likelihood function to discount the information from previous visits, and the posterior distribution of the five-category random variable is obtained using the Dirichlet prior for **p** and a Beta (1, 1) as a power prior for .

```
Visit 1: [\mathbf{p} | \mathbf{n}_1] \propto [\mathbf{n}_1 | \mathbf{p}]^*[\mathbf{p}]

Visit 2: [\mathbf{p}, a_0 | \mathbf{n}_1, \mathbf{n}_2] \propto [\mathbf{n}_2 | \mathbf{p}]^*[\mathbf{n}_1 | \mathbf{p}]^{a_0} *[\mathbf{p}]^*[a_0]

......

Visit M: [\mathbf{p}, a_0 | \mathbf{n}_1, \mathbf{n}_2, ..., \mathbf{n}_M] \propto [\mathbf{n}_M | \mathbf{p}]^*[\mathbf{n}_1, \mathbf{n}_2, ..., \mathbf{n}_{M-1} | \mathbf{p}]^{a_0} *[\mathbf{p}]^*[a_0]
```



MODEL FIT

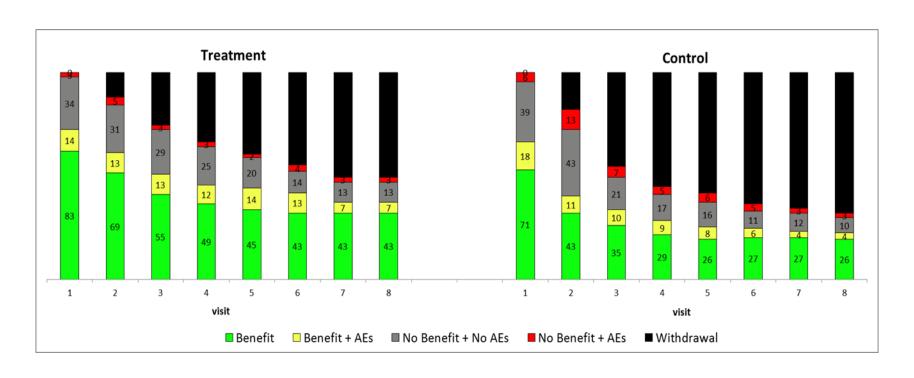
 The model fit of the two models (with and without power prior) is assessed through the conditional predictive ordinate (CPO) and the logarithm of the pseudo-marginal likelihood (LPML). The larger the value of LPML, the better fit the model is. Here, n_(i) is the data with n_i removed.

$$CPO_{i} = f(\mathbf{n}_{i} | \mathbf{n}_{(i)}) = \int [\mathbf{n}_{i} | \mathbf{p}, \mathbf{n}_{(i)}][p | \mathbf{n}_{(i)}]d\mathbf{p}$$

$$LPML = \sum_{i=1}^{5} \log f(\mathbf{n}_{i} | \mathbf{n}_{(i)}) = \sum_{i=1}^{5} \log CPO_{i}$$



BACK TO OUR EXAMPLE 1: HYDROMORPHONE

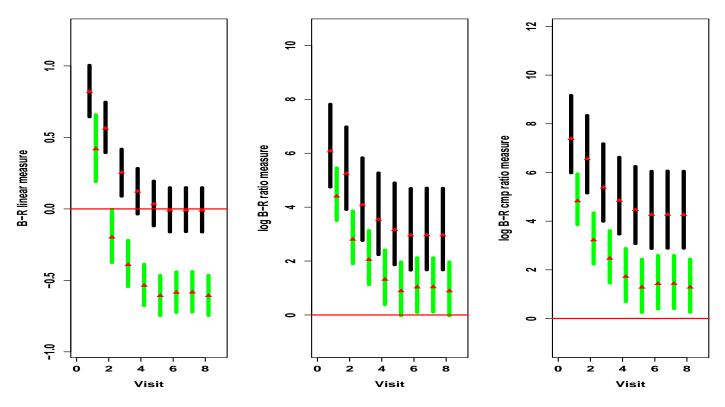




AGGREGATED BR MEASURES



 Posterior means and posterior 95% credible intervals for BR-L, log(BR-R) and log(BR-CR) averaged over all subjects.



Red triangles indicates posterior means, black and green bars are credible intervals from treatment and control arms respectively.



TESTING EQUALITY OF BENEFIT-RISK MEASUREMENTS BETWEEN TREATMENT AND CONTROL ARMS

Table: 95% credible intervals of ABRD-L, ABRD-R and ABRD-CR at each visit.

	Visits								
	1	2	3	4	5	6	7	8	
ABRD-L	(0.103,	(0.499,	(0.409,	(0.438,	(0.431,	(0.363,	(0.365,	(0.386,	
	0.695)	1.019)	0.874)	0.874)	0.857)	0.788)	0.787)	0.809)	
ABRD-R	(-0.049,	(0.726,	(0.307,	(0.506,	(0.554,	(0.207,	(0.218,	(0.362,	
	3.612)	4.392)	3.981)	4.168)	4.229)	3.879)	3.883)	4.008)	
ABRD-CR	(0.753,	(1.505,	(1.085,	(1.298,	(1.359,	(1.007,	(1.000,	(1.153,	
	4.594)	5.380)	4.975)	5.188)	5.254)	4.892)	4.889)	5.056)	

There is significant evidence of increasing benefit-risk measurements on the treatment arm than the control arm.



PERSONALIZED BR ASSESSMENT: MOTIVATION

- Zhao et al. (2014) method
 - Works on the aggregated level
 - Ignore the subject-level response differences
 - Assume subjects are exchangeable at each visit
- The proposed model:
 - Incorporates subject-level effects
 - Uses Dirichlet process model to accommodate the similarity of subject responses
 - Provides subject-specific response probabilities and benefit-risk analysis

PERSONALIZED BR BAYESIAN MODEL



- A generalized linear model (McCullagh and Nelder, 1989)
 is used to define the log-odds for all categories with
 respect to the reference (the first category).
- Define the log-odds of the i-th subject's response falling into category j ($y_{ik} = j$) with respect to category 1, at visit k as

$$\eta_{ijk} = \log \frac{\Pr(y_{ik} = j)}{\Pr(y_{ik} = 1)} = \beta_{ij} + \gamma_k$$

 $\beta_i = (\beta_{i2}, \dots, \beta_{iJ})$: subject-level category effects $(i = 1, \dots, N)$ γ_k $(k = 1, \dots, K)$: random longitudinal visit effects



WHY USE DIRICHLET PROCESS?

- The proposed model considers subject-level differences
- There also exists similarity among the subject responses (clusters for subjects)
 - Within cluster, subject-level effects are same
 - Across clusters, subject-level effects are distinct
- Mixture density for subject-level differences
 - Unknown number of Clusters

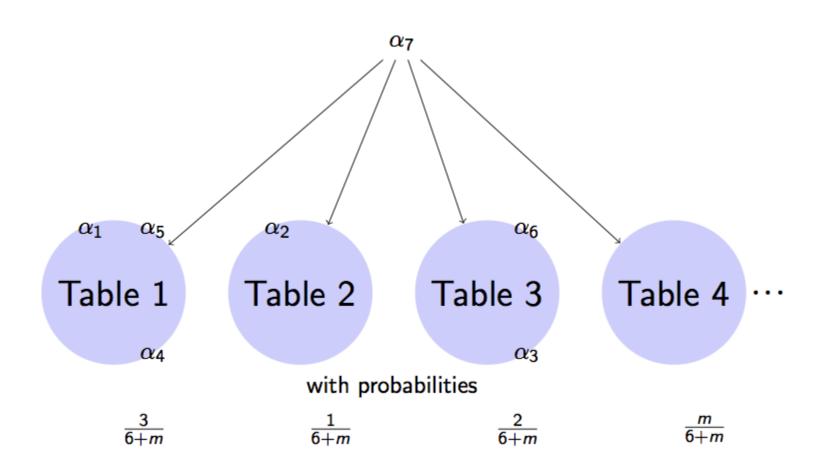


PRIOR SPECIFICATION

- Dirichlet process prior for the subject-level effects β_{ij} .
 - $\cdot \beta_{ij} \sim F, F \sim DP(m, F_0)$
 - $m \ge 0$: the concentration parameter
 - · F_0 : baseline distribution $N(0, \Sigma_{\beta})$
- Visit effects $\gamma = (\gamma_1, ..., \gamma_K)^{'} \sim N(0, \Sigma_{\gamma})$

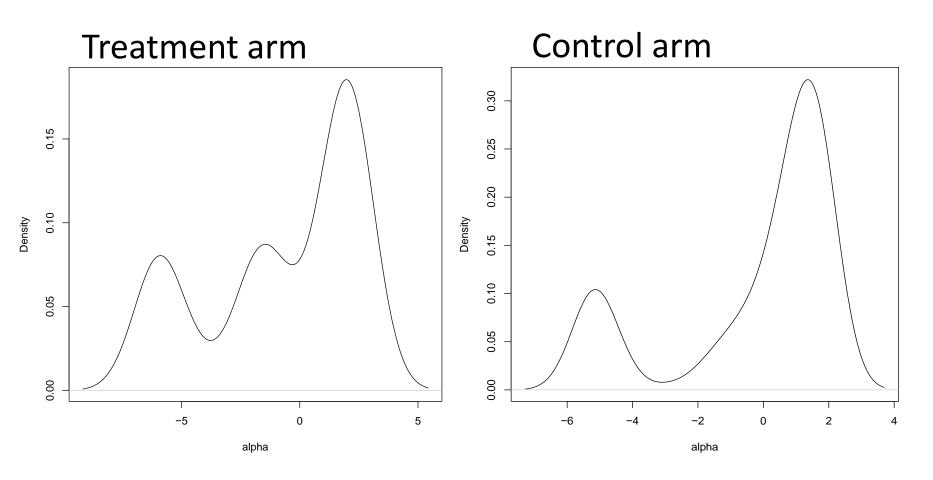


DP CLUSTERS: CHINESE RESTAURANT PROCESS





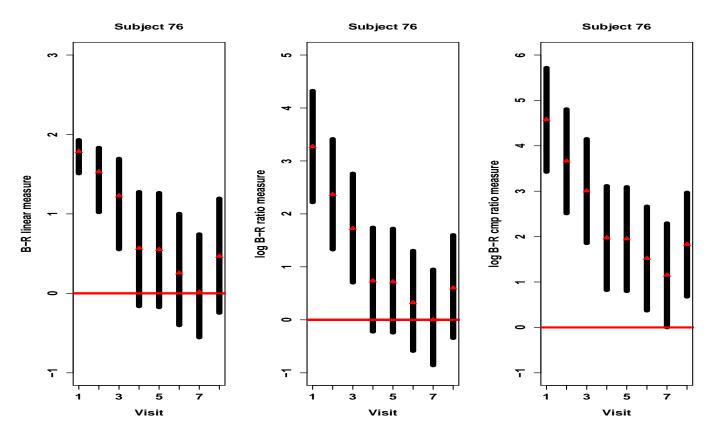
CLUSTERED SUBJECT-LEVEL EFFECTS





PERSONALIZED BENEFIT-RISK MEASURES

Posterior profiles of three benefit-risk measures for subject, i = 76 from treatment arm. Vertical bars indicate 95% credible intervals of measurements, red triangles are corresponding posterior means.





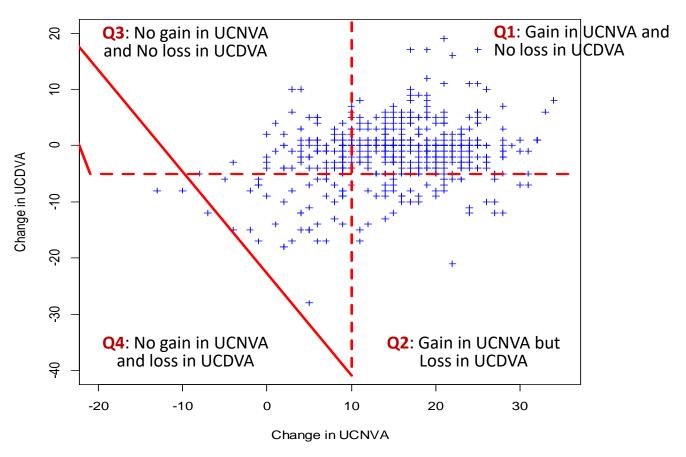
EXAMPLE 2: OPHTHALMIC DEVICE

- Benefit is measured by improvement in uncorrected (without glasses or contact lenses) near or intermediate visual acuity (UCNVA).
- Risk is measured by uncorrected distance visual acuity (UCDVA), since subjects may be giving up distance vision for some gain in near or intermediate vision.
- The data on the joint evaluation of benefit and risk,
 (UCNVA, UCDVA), of the device is given on the next slide





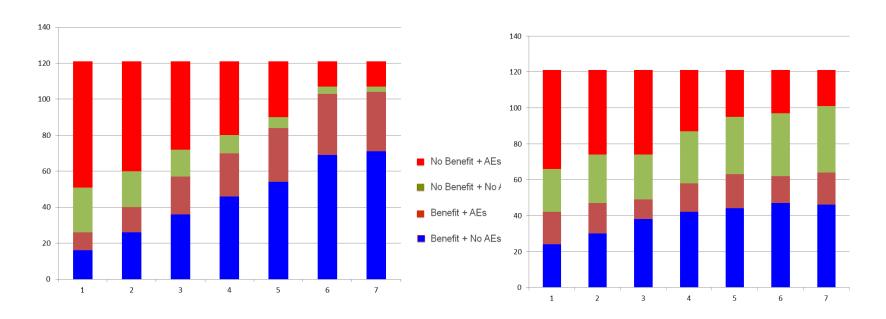
Change in UCNVA vs Change in UCDVA from Baseline





EXAMPLE 2: DATA FROM OPHTHALMIC STUDY: HYPOTHETICAL

 Two-arm study, 242 patients were randomly assigned to treatment and control arms. Patients were followed up for seven visits, and their outcomes were assigned to one of the four benefit-risk categories.





BENEFIT-RISK (BR) MEASURES

Interested in three BR measures:

$$(Q_1 + Q_2) - (Q_3 + Q_4) \rightarrow$$
 Difference between Benefit and No-Benefit $(Q_1 + Q_2)/(Q_3 + Q_4) \rightarrow$ Ratio of Benefit to No-Benefit $Q_1/Q_4 \rightarrow$ Ratio of Best Quadrant to Worst Quadrant

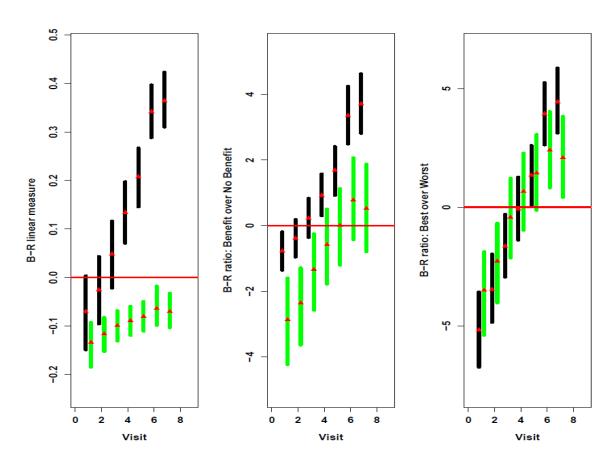
BR - L =
$$\sum_{i=1}^{2} w_{i} p_{i,T} - \sum_{i=3}^{4} w_{i} p_{i,T} : (Q_{1} + Q_{2}) - (Q_{3} + Q_{4})$$

BR - R1 = $\frac{\sum_{i=1}^{2} w_{i} p_{i,T}}{\sum_{i=3}^{4} w_{i} p_{i,T}} : (Q_{1} + Q_{2}) / (Q_{3} + Q_{4})$
BR - R2 = $\frac{w_{1} p_{1,T}}{w_{4} p_{4,T}} : Q_{1} / Q_{4}$





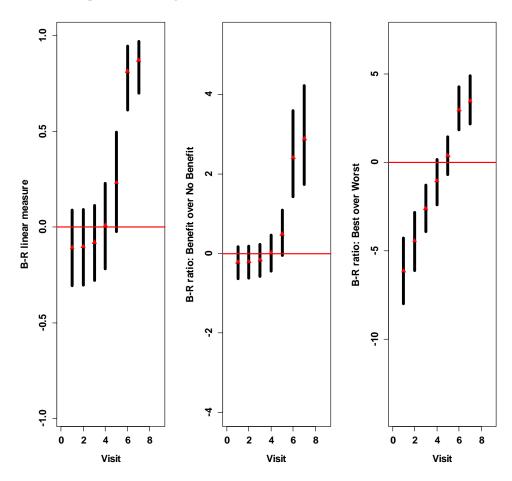
Posterior 95% CIs of BR-L, BR-R1, and BR-R2 at visits k = 1, . . , 7 for treatment arm (black) and control arm (green)



PERSONALIZED BENEFIT-RISK ASSESSMENT



Posterior 95% CIs of BR-L, BR-R1, and BR-R2 at visits k = 1, . . , 7 for one single subject i = 66, from the treatment arm





CONCLUSION

- Quantitative measures of benefit and risk are important aspect in the drug and device evaluation process
- The Bayesian method is a natural method for longitudinal data by sequentially updating the prior; Power prior can be used to discount information from previous visits.
- Use of Dirichlet process for modeling subject-level longitudinal BR categorical data seems to be a reasonable approach
- Model selection approaches can be incorporated to compare model fits
- More quantitative work on BR assessment is warranted

Other References

FDA

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THANK YOU!